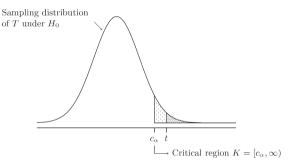
Statistical Methods for Data Science Lesson 22 - Multiple comparisons. Fitting distributions.

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Critical values and p-values



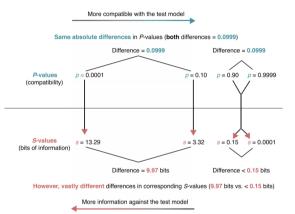
- Critical region K: the set of values that reject H_0 in favor of H_1 at significance level α
- Critical values: values on the boundary of the critical region
- *p-value*: the probability of obtaining test results at least as extreme as the results actually observed, under the assumption that H_0 is true
- $t \in K$ iff *p*-value $\leq \alpha$

Misues of *p*-values

Misinterpretations of p-values, Greenland et al., 2016

- The p-value is the probability that the null hypothesis is true, or the probability that the alternative hypothesis is false. A p-value indicates the degree of compatibility between a dataset and a particular hypothetical explanation
- The 0.05 significance level is the one to be used: No, it is merely a convention. There is no reason to consider results on opposite sides of any threshold as qualitatively different.
- A large p-value is evidence in favor of the test hypothesis: A p-value cannot be said to favor the test hypothesis except in relation to those hypotheses with smaller p-values
- If you reject the test hypothesis because p ≤ 0.05, the chance you are in error is 5%: No, the chance is either 100% or 0%. The 5% refers only to how often you would reject it, and therefore be in error.

s-values



- Shannon information value or surprisal value (s-value) is $-\log_2 p$
- $p = 0.05 \Rightarrow s = 4.3$ no more surprising than getting all heads on 4 fair coin tosses.
- $p = 0.005 \Rightarrow p = 7.64$ no more surprising than getting all heads on 8 fair coin tosses.

The multiple comparisons problem

• Single test $H_0: \theta = 0$, with significance level $\alpha = 0.05$

[false positive rate]

- test is called *significant* when we reject H_0
- α is Type I error, probability of rejecting H_0 when it is true
- Multiple tests, say m = 20
 - ► E.g., $H_0^i : \theta_i = 0$ for i = 1, ..., m where θ_i is the expectation of a subpopulation
- What is the probability of rejecting at least one H_0^i when all of them are true?

$$\mathsf{P}(ext{at least one reject}) = \mathsf{P}(\cup_{i=1}^m \{ \mathsf{p}_i \leq lpha \}) = 1 - \mathsf{P}(\cap_{i=1}^m \{ \mathsf{p}_i > lpha \}) = 1 - (1 - lpha)^m$$

and then $1-(0.95)^{20}\approx 0.64$

Family-wise error rate (FWER)

The FWER is the probability of making at least one Type I error in a family of n tests. If the tests are independent:

$$\alpha_{FWER} = 1 - (1 - \alpha)^m$$

If the test are dependent: $\alpha_{FWER} \leq m \cdot \alpha$

Multiple comparisons: corrections

• Bonferroni correction (most conservative one):

$$\alpha = \frac{\alpha_{FWER}}{m}$$

Hence, $p < \alpha$ iff $p \cdot m < \alpha_{FWER}$

• *Šidák correction* (exact for independent tests):

$$lpha = 1 - (1 - lpha_{FWER})^{1/m}$$

Hence, $p < \alpha$ iff $1 - (1 - p)^m < \alpha_{FWER}$ See R script

False Discovery Rate and *q*-values

| | | True state of nature | |
|--|------------------|----------------------|----------------|
| | | H_0 is true | H_1 is true |
| Our decision on the basis of the data | Reject H_0 | False Positive | True Positive |
| | Not reject H_0 | True Negative | False Negative |

- False Positive Rate: FPR = FP/(FP + TN)
 - Corrections control for *FPR* since $FWER = P(FP > 0|H_0^i \ i = 1, ..., m)$
- Drawback: acting on α increases FNR = FN/(FN + TP)
- False Discovery Rate: FDR = FP/(FP + TP)
 - FDR = 0.05 means 5% of rejected H_0 's are actually true
- *q*-value is $P(H_0|T \ge t)$
 - *FDR* can be controlled by requiring $q \leq$ threshold

See R script

 $[p = P(T \ge t | H_0)]$

Distribution fitting

- Dataset x_1, \ldots, x_n realization of $X_1, \ldots, X_n \sim F$
- What is a plausible F?
- Parametric approaches:
 - Assume $F = F(\lambda)$ for some family F, and estimate λ as $\hat{\lambda}$
 - □ Maximum Likelihood Estimation (point estimate):

$$\hat{\lambda} = {\sf argmax}_{\lambda} {\sf L}(\lambda)$$

□ Parametric bootstrap (*p*-value):

$$T_{ks} = \sup_{a \in \mathbb{R}} |F_n^*(a) - F_{\hat{\Lambda}^*}(a)|$$

- Non-parametric approaches:
 - Empirical distribution
 - Kernel Density Estimation
- Goodness of fit: how good is F in fitting the data?

Goodness of fit

- Loss functions (to be minimized)
 - ► Akaike information criterion (AIC), balances model fit against model simplicity

$$AIC(F(\lambda)) = 2|\lambda| - 2\ell(\lambda)$$

Bayesian information criterion (BIC), stronger balances over model simplicity

$$BIC(F(\lambda)) = |\lambda| \log n - 2\ell(\lambda)$$

- Statistics (continuous data):
 - ▶ KS test $H_0: X \sim F$ $H_1: X \not\sim F$ with Kolmogorov-Smirnov (KS) statistic:

$$D = \sup_{a \in \mathbb{R}} |F_n(a) - F(a)| \sim K$$

▶ LR test $H_0: X \sim F_1$ $H_1: X \sim F_2$ with the likelihood-ratio test:

$$\lambda_{LR} = \log \frac{L(F_1(\lambda_1))}{L(F_2(\lambda_2))} = \ell(F_1(\lambda_1)) - \ell(F_2(\lambda_2)) \quad \text{with} \quad -2\lambda_{LR} \sim \chi^2$$

See R script

Goodness of fit

Chi-square distribution

The Chi-square distribution with k degrees of freedom $\chi^2(k)$ has density:

$$f(x) = \frac{1}{2^{k/2} \Gamma(k/2)} x^{k/2-1} e^{-x/2}$$

Let
$$X_1, ..., X_k \sim N(0, 1)$$
. Then $Y = \sum_{i=1}^k X_i^2 \sim \chi^2(k)$

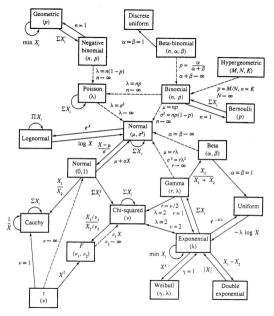
- Statistics (discrete data):
 - ▶ Pearson's Chi-Square test H_0 : $X \sim F$ H_1 : $X \not\sim F$ with χ^2 statistic:

$$\chi^{2} = \sum_{N_{i} > 0} \frac{(N_{i} - n_{i})^{2}}{n_{i}} = n \cdot \sum_{N_{i} > 0} \frac{(N_{i} / n - p(i))^{2}}{p(i)} \sim \chi^{2}(df)$$

where N_i number of observations of value i, $n_i = n \cdot p(i)$ expected number of observations (rescaled), and $df = |\{i \mid N_i > 0\}| - 1$ is the number of observed values minus 1. $\chi^2 = \infty$ if for some i: $n_i = 0$

See R script

Common distributions



Relationships among common distributions. Solid lines represent transformations and special cases, dashed lines represent limits. Adapted from Leemis (1986).

Comparing two datasets

- Dataset x_1, \ldots, x_n realization of $X_1, \ldots, X_n \sim F_1$
- Dataset y_1, \ldots, y_m realization of $Y_1, \ldots, Y_n \sim F_2$
- $H_0: F_1 = F_2$ $H_1: F_1 \neq F_2$
- Continuous data: KS statistics

$$D = \sup_{a \in \mathbb{R}} |F_1(a) - F_2(a)| \sim K$$

• Discrete data: χ^2 statistics

$$\chi^2 = \sum_{R_i > 0 \lor S_i > 0} \frac{\left(\sqrt{\frac{m}{n}}R_i - \sqrt{\frac{n}{m}}S_i\right)^2}{R_i + S_i} \sim \chi^2(df)$$

where R_i (resp., S_i) is the number of observations in x_1, \ldots, x_n (resp., y_1, \ldots, y_m) of value i, $df = |\{i | R_i > 0 \lor S_i > 0\}| - 1$

See R script